

**EX:** (This problem is motivated by problem of using the `rand()` function in Matlab<sup>®</sup> to create arbitrary probability density functions.) Given three independent random variables,  $V$ ,  $W$ , and  $Z$ , that are uniformly distributed on  $[0, 1]$ , describe a step-by-step calculation that yields random variables  $X$  and  $Y$  with the following joint density function (whose footprint is shaped like a diamond centered on the origin):

$$f(x, y) = \begin{cases} \frac{1}{2} & |X| + |Y| \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

Hint: First generate  $X$  from the density function  $f_X(x)$  using some simple algebra involving  $V$  and  $W$ . Then generate  $Y$  from the conditional probability density function  $f(y | X)$ . Use  $Z$  and some simple algebra to create  $Y$ .

**SOL'N:** The plot below shows the support (or footprint) of  $f(x, y)$ .

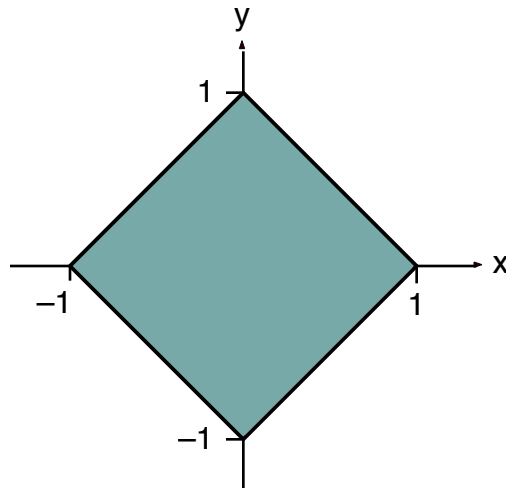


Fig. 1. Support (or footprint) of  $f(x, y)$ .

In a 3-dimensional view, the diamond shape of  $f(x, y)$  has a constant height of  $1/2$ .

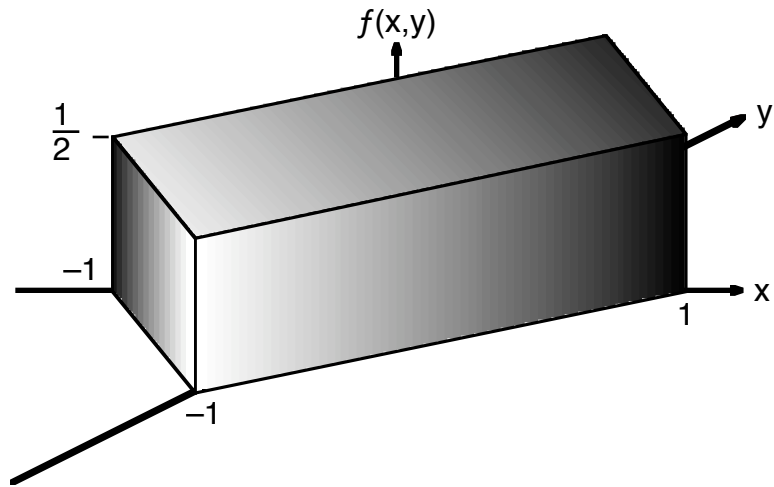


Fig. 2. 3-dimensional plot of  $f(x, y)$ .

The rationale for the hint is that we can write the joint probability,  $f(x, y)$ , as the product of a density function for  $x$  alone and a conditional probability for  $y$  given  $x$ :

$$f(x, y) = f(y | X = x)f_X(x)$$

This means that we can first pick  $X$  distributed as  $f_X(x)$  and then pick  $Y$  distributed as  $f(y | X)$ .

To find  $f_X(x)$ , we use the standard formula for integration in the  $y$  direction:

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

Fig. 3, below, shows the limits of the integral for a particular value of  $x = x_0$  as the endpoints of a cross-section in the  $y$  direction. The value of  $f(x, y)$  over this segment is one-half.

$$f_X(x_0) = \int_{-(1-|x_0|)}^{1-|x_0|} \frac{1}{2} dy = 1 - |x_0|$$

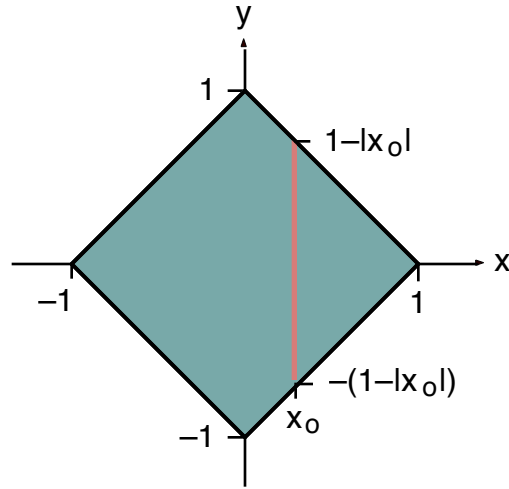


Fig. 3. Top view of cross-section used to calculate  $f_X(x_0)$  and  $f(y | X = x_0)$ .

This above formula, written using absolute value, actually holds for any positive or negative value of  $x_0$ , and we have the following formula for probability density of  $X$ :

$$f_X(x) = 1 - |x|$$

Fig. 4 shows that  $f_X(x)$  is triangular.

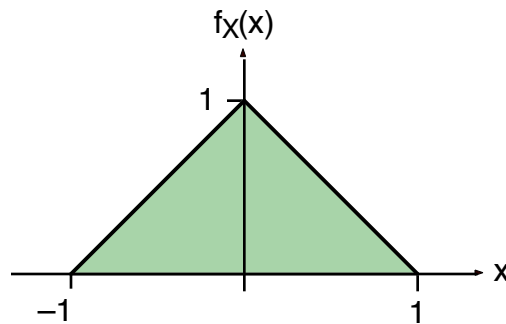


Fig. 4. Plot of  $f_X(x)$ .

There are two straightforward ways to generate a random variable,  $X$ , with this probability density function. The first is to add two uniformly

distributed random variables together (and subtract one to give a mean of zero):

$$X = V + W - 1$$

The probability density function for  $X$  is computed as a convolution integral. We start with the probability density of  $V$  and find the probability density that  $W = X - (V - 1)$ . We integrate this product over possible values of  $V$ .

$$f_X(x) = \int_0^1 f_V(v) f_W(w = x - (v - 1)) dv$$

We observe that  $f_W(w) = 1$  when  $0 < w < 1$ .

$$f_X(x) = \int_0^1 f_V(v) \cdot \begin{cases} 1 & 0 < x - (v - 1) < 1 \\ 0 & \text{otherwise} \end{cases} dv$$

Rearranging the inequality to express it in terms of  $v$  yields the following expression:

$$f_X(x) = \int_0^1 f_V(v) \cdot \begin{cases} 1 & x < v < x + 1 \\ 0 & \text{otherwise} \end{cases} dv$$

Substituting  $f_V(v) = 1$  and translating the expression for  $f_W(w)$  into modifications of the limits of integration yields the following expression for the density function shown in Fig. 4:

$$f_X(x) = \int_{\max(0,x)}^{\min(1,x+1)} 1 \, dv = \begin{cases} \int_0^{x+1} 1 \, dv = x + 1 & -1 < x < 0 \\ \int_x^1 1 \, dv = 1 - x & 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$$

From the above discussion, the step-by-step procedure for calculating  $X$  is to use the following simple formula:

$$X = V + W - 1$$

Another way to obtain a random variable with the density function shown in Fig. 4 is to transform a single uniform random variable such as  $V$  by matching the cumulative distribution functions of  $X$  and  $V$ .

The cumulative distribution function for  $V$  is easily computed:

$$F_V(v) = \int_{-\infty}^v f_V(v)dv = \begin{cases} 0 & v < 0 \\ v & 0 < v < 1 \\ 1 & v > 1 \end{cases}$$

The cumulative distribution function for  $X$  is quadratic since  $f_X(x)$  is linear.

$$F_X(x) = \int_{-\infty}^{\infty} f_X(x)dx = \begin{cases} 0 & x < -1 \\ \frac{1}{2}(x+1)^2 & -1 < x < 0 \\ 1 - \frac{1}{2}(x-1)^2 & 0 < x < 1 \\ 1 & x > 1 \end{cases}$$

Given a value for  $V$ , we find a value of  $X$  such that  $F_X(x) = F_V(V)$ . This translates in the following equation:

$$X \text{ satisfies } \begin{cases} \frac{1}{2}(X+1)^2 = V & 0 < V < \frac{1}{2} \\ 1 - \frac{1}{2}(X-1)^2 = V & \frac{1}{2} < V < 1 \end{cases}$$

or

$$X = \begin{cases} X = \sqrt{2V} - 1 & 0 < V < \frac{1}{2} \\ X = \sqrt{2(1-V)} + 1 & \frac{1}{2} < V < 1 \end{cases}$$

Now that we have  $X$ , we use the conditional probability density function,  $f(y | X)$  for  $Y$ . We find  $f(y | X)$  by first taking a cross section of  $f(x, y)$  at  $x = X$ , as shown in Fig. 5.

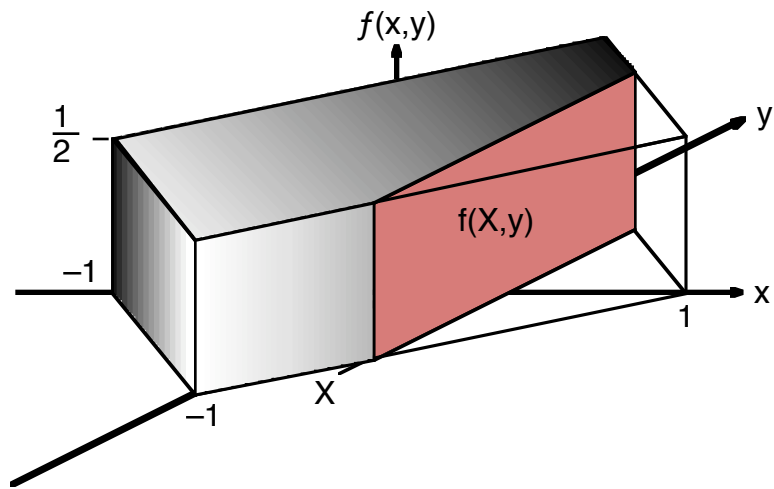


Fig. 5. Cross-section used to calculate  $f_X(X)$  and  $f(y | X)$ .

We scale the cross section vertically so it will have a total area equal to one. Fig. 6 shows the result.

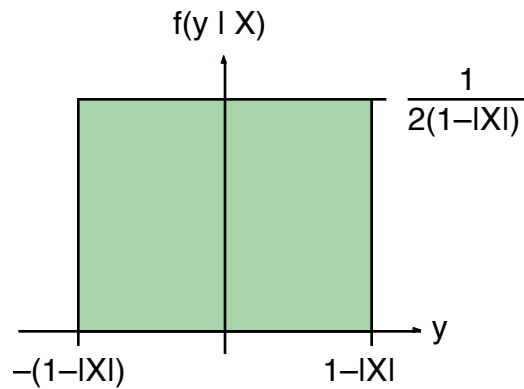


Fig. 6. Conditional probability  $f(y | X)$ .

We obtain this distribution by shifting and scaling a (0,1) uniform distribution such as  $Z$ .

$$Y = 2(1 - |X|) \left( Z - \frac{1}{2} \right)$$