

TOOL: The variables Z and χ^2 , used in the derivation of the t -distribution are independent.

$$T = \frac{\bar{X} - \mu}{S/\sqrt{n}}$$

The relevant definitions are as follows:

n = number of data points, X_i , (which are independent and normally distributed)

$$v = n - 1$$

$$\bar{X} \equiv \frac{1}{n} \sum_{i=1}^n X_i \equiv \text{sample mean}$$

$$S^2 \equiv \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 \equiv \text{sample variance}$$

$$Z \equiv \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \quad \text{and} \quad \chi^2 \equiv v \frac{S^2}{\sigma^2}$$

PROOF: For convenience, we define the following standard normal random variables:

$$Z_i \equiv \frac{X_i - \mu}{\sigma}$$

We may define Z in terms of the Z_i :

$$Z \equiv \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} = \sqrt{n} \frac{1}{n} \sum_{i=1}^n \frac{X_i - \mu}{\sigma} = \frac{1}{\sqrt{n}} \sum_{i=1}^n Z_i$$

We may also define χ^2 in terms of the Z_i :

$$\chi^2 = \frac{n-1}{\sigma^2} \cdot \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = \sum_{i=1}^n \left(\frac{X_i - \mu}{\sigma} - \frac{\bar{X} - \mu}{\sigma} \right)^2$$

or

$$\chi^2 = \sum_{i=1}^n \left(Z_i - \frac{1}{n} \sum_{j=1}^n Z_j \right)^2$$

To prove the independence of Z and χ^2 , we show that $E(Z \cdot \chi^2) = E(Z)E(\chi^2)$.

Since $E(Z) = 0$, however, this means we will show that $E(Z \cdot \chi^2) = 0$.

We employ our definitions of Z and χ^2 in terms Z_i 's at the outset:

$$E(Z \cdot \chi^2) = E \left(\sqrt{n} \frac{1}{n} \sum_{i=1}^n Z_i \cdot \sum_{i=1}^n \left(Z_i - \frac{1}{n} \sum_{j=1}^n Z_j \right)^2 \right)$$

We may rewrite the second summation as follows, (see [1], p. 233):

$$E(Z \cdot \chi^2) = \frac{1}{\sqrt{n}} E \left(\sum_{i=1}^n Z_i \cdot \left[\sum_{i=1}^n Z_i^2 - \frac{1}{n} \left(\sum_{i=1}^n Z_i \right)^2 \right] \right)$$

Multiplying by n/n and rearranging slightly yields the following expression:

$$E(Z \cdot \chi^2) = \frac{1}{n\sqrt{n}} E \left(n \sum_{i=1}^n Z_i^2 \sum_{i=1}^n Z_i - \left(\sum_{i=1}^n Z_i \right)^3 \right)$$

The Z_i are independent with mean zero, allowing us to eliminate terms that involve products of different Z 's. In other words, the only terms that will contribute to the expected value will be those involving a Z_i that is cubed.

$$E(Z \cdot \chi^2) = \frac{1}{n\sqrt{n}} E \left(n \sum_{i=1}^n Z_i^3 - \sum_{i=1}^n Z_i^3 \right) = \frac{1}{n\sqrt{n}} E \left((n-1) \sum_{i=1}^n Z_i^3 \right)$$

$$E(Z \cdot \chi^2) = \frac{n-1}{n\sqrt{n}} E \left(\sum_{i=1}^n Z_i^3 \right) = \frac{n-1}{n\sqrt{n}} n E(Z_1^3) = \frac{n-1}{\sqrt{n}} E(Z_1^3)$$

Since a standard normal distribution is symmetric around zero, the expected value of each cubed term is zero, however.

$$E(Z \cdot \chi^2) = 0$$

It follows that Z and χ^2 are independent.

REF: [1] Ronald E. Walpole, Raymond H. Myers, Sharon L. Myers, and Keying Ye, *Probability and Statistics for Engineers and Scientists*, 8th Ed., Upper Saddle River, NJ: Prentice Hall, 2007.