

# ① Unbiased estimator

$$a) E(\hat{\theta}) = E(a^T x)$$

$$E(\hat{\theta}) = a^T E(x)$$

$$E(\hat{\theta}) = a^T \theta s = \theta a^T s = \theta$$

$$b) a^T C a = a^T E[ww^T] a$$

$$= a^T E[(x - E(x))(x - E(x))^T] a$$

$$= E[a^T (x - E(x))(x - E(x))^T a]$$

$$= E\left[\underbrace{(a^T x - a^T E(x))}_{\hat{\theta}} \underbrace{(a^T x - a^T E(x))}_{E(\hat{\theta})}\right]$$

$$= E\left[(\hat{\theta} - E(\hat{\theta}))(\hat{\theta} - E(\hat{\theta}))^T\right]$$

$$= \text{Var}(\hat{\theta})$$

Note: The quantity inside the expectation is a scalar; hence we say  $\text{var}(\hat{\theta})$  instead of  $C_{\hat{\theta}}$ .

$$\textcircled{2} \quad a) \quad \ln p(y; \alpha) = \ln \frac{N!}{k!(N-k)!}$$

$$+ (N-k) \ln \alpha$$

$$+ k \ln (1-\alpha)$$

To find MLE for  $\alpha$ :

$$\frac{\partial \ln p(y; \alpha)}{\partial \alpha} = \frac{N-k}{\alpha} - \frac{k}{1-\alpha} = 0$$

$$\frac{(N-k)(1-\alpha) - k\alpha}{\alpha(1-\alpha)} = 0$$

~~scribbled out text~~

$$(N-k)(1-\alpha) - k\alpha = 0$$

$$N - N\alpha - k + k\alpha - k\alpha = 0$$

$$N - k = N\alpha \quad \Rightarrow \quad \hat{\alpha} = \frac{N-k}{N}$$

b) This is a mixture of 2 Gaussians:

$$p(x[n]; \alpha) = \frac{\alpha}{\sqrt{2\pi A}} \exp\left[-\frac{x[n]^2}{2A}\right] + \frac{1-\alpha}{\sqrt{2\pi B}} \exp\left[-\frac{x[n]^2}{2B}\right]$$

$$\text{Var}(x[n]) = \alpha A + (1-\alpha) B$$

③ a) Since the MLE for the transformation of a parameter is the transformation of the MLE for the parameter:

$$\hat{\alpha}_{MLE} = \frac{(\hat{A}_{MLE})^2}{\sigma^2}$$

Now we know  $\hat{A}_{MLE} = \frac{1}{N} \sum_{n=0}^{N-1} x[n]$

$$\text{so } \hat{\alpha}_{MLE} = \frac{\left(\frac{1}{N} \sum_{n=0}^{N-1} x[n]\right)^2}{\sigma^2}$$

$$b) E[\hat{\alpha}_{MLE}] = \frac{1}{\sigma^2} E[(\hat{A}_{MLE})^2]$$

$$= \frac{1}{\sigma^2} \left( E[\hat{A}_{MLE}]^2 + \text{var}(\hat{A}_{MLE}) \right)$$

$$= \frac{1}{\sigma^2} \left( A^2 + \frac{\sigma^2}{N} \right) = \frac{A^2}{\sigma^2} + \frac{1}{N}$$

$$\neq \frac{A^2}{\sigma^2}$$

$$c) \hat{\alpha}_{MLE} \stackrel{a}{\sim} N\left(\frac{A^2}{\sigma^2}, I(\alpha)^{-1}\right)$$

$$\text{Var}(\hat{\alpha}_{MLE}) = J(\alpha)^{-1} = \frac{(\partial \alpha / \partial A)^2}{-E\left[\frac{\partial^2 p(x; A)}{\partial A^2}\right]} = \left(\frac{\partial \alpha}{\partial A}\right)^2 I(A)^{-1}$$

asymptotic

$$= \left(\frac{2A}{\sigma^2}\right)^2 \frac{\sigma^2}{N} = \frac{4A^2}{\sigma^2 N}$$

④ a) For an orthonormal basis,  $H^T H = I$   
 therefore  $\hat{\theta} = H^T x$ . For the Haar basis  
 with  $N=4$

$$\begin{pmatrix} \hat{\theta}_1 \\ \hat{\theta}_2 \end{pmatrix} = \begin{pmatrix} h_1^T \\ h_2^T \end{pmatrix} x = \frac{1}{2} \begin{pmatrix} 1 & 1 & 1 & 1 \\ 1 & 1 & -1 & -1 \end{pmatrix} \begin{pmatrix} x[0] \\ x[1] \\ x[2] \\ x[3] \end{pmatrix}$$

$$\text{so } \hat{\theta}_1 = \frac{1}{2} \sum_{n=0}^3 x[n]$$

$$\hat{\theta}_2 = \frac{1}{2} (x[0] + x[1] - x[2] - x[3])$$

$$\begin{aligned} \text{b) } \varepsilon &= \|x - H\hat{\theta}\|^2 = \|x - HH^T x\|^2 = \|(I - HH^T)x\|^2 \\ &= x^T (I - HH^T)(I - HH^T)x \\ &= x^T [I - HH^T - HH^T + \underbrace{HH^T HH^T}_{\hat{I}}] x \\ &= x^T [I - 2HH^T + \hat{I}] x \\ &= x^T (I - HH^T) x \end{aligned}$$