

$$\textcircled{1} \quad x[n] = R + w[n] \quad n=0, \dots, N-1$$

$w[n] \sim N(0, 1)$ WGN, w & R independent

$R \sim N(100, 0.011)$ prior PDF

$$\text{So } \mu_R = 100 \quad \sigma_R^2 = 0.011 \quad \sigma^2 = 1$$

We know in this case $\text{Bmse}(\hat{R}) = \frac{1}{N/\sigma^2 + 1/\sigma_R^2}$
(Example 10-1)

$$\text{We need } \frac{1}{N/\sigma^2 + 1/\sigma_R^2} \leq 0.1^2$$

$$\frac{\sigma^2 \sigma_R^2}{N\sigma_R^2 + \sigma^2} \leq 0.01$$

$$\text{So } \cancel{0.1} \times 0.011 \leq \cancel{0.01} (N \times 0.011 + 0.1)$$

$$\frac{0.11 - 0.1}{0.011} \leq N$$

$9.09 \leq N$ But N has to be an integer so we need $N \geq 10$

$$\textcircled{2} \quad x[n] = Ar^n + w[n] \quad n=0, \dots, N-1$$

r known, w WGN with σ^2 , $A \sim N(0, \sigma_A^2)$

$$\underline{x} = \underbrace{\begin{bmatrix} 1 \\ r \\ r^2 \\ \vdots \\ r^{N-1} \end{bmatrix}}_H A + \underline{w}$$

Notice $\mu_A = 0$
 $C_A = \sigma_A^2$
 $C_w = \sigma^2 I$

$$\hat{A}_{\text{MMSE}} = E[A|x] = \mu_A + (C_A^{-1} + H^T C_w^{-1} H)^{-1} H^T C_w^{-1} (x - H\mu_A)$$

$$\text{So } \hat{A}_{\text{MMSE}} = 0 + \left(\frac{1}{\sigma_A^2} + \frac{1}{\sigma^2} H^T H \right)^{-1} H^T \frac{1}{\sigma^2} (x - H_0)$$

$$H^T H = \sum_{n=0}^{N-1} r^{2n} \quad \text{notice this is a scalar}$$

$$H^T x = \sum_{n=0}^{N-1} r^n x[n] \quad //$$

$$\text{Therefore } \hat{A}_{\text{MMSE}} = \frac{\frac{1}{\sigma^2} \sum_{n=0}^{N-1} r^n x[n]}{\frac{1}{\sigma_A^2} + \frac{1}{\sigma^2} \sum_{n=0}^{N-1} r^{2n}}$$

$$\begin{aligned} \text{Bmse}(\hat{A}) &= C_{A|X} = \left(C_A^{-1} + H^T C_w H \right)^{-1} \\ &= \left(\frac{1}{\sigma_A^2} + \frac{1}{\sigma^2} \sum_{n=0}^{N-1} r^{2n} \right)^{-1} \end{aligned}$$

$$\textcircled{3} \quad p(x, y) = \frac{1}{2\pi \sqrt{|C|}} \exp \left[- \frac{[x \ y] C^{-1} \begin{bmatrix} x \\ y \end{bmatrix}}{2} \right]$$

since x & y 0 mean

when $x = x_0$ we have

$$g(y) = p(x_0, y) = \frac{1}{2\pi |C|^{1/2}} e^{-h(y)/2}$$

$$\begin{aligned} \text{where } h(y) &= [x_0 \ y] \begin{bmatrix} 1 & -\rho \\ -\rho & 1 \end{bmatrix} \begin{bmatrix} x_0 \\ y \end{bmatrix} / (1 - \rho^2) \\ &= \frac{x_0^2 + y^2 - 2\rho x_0 y}{1 - \rho^2} \end{aligned}$$

To maximize $g(y)$ minimize $h(y)$

$$\frac{\partial h}{\partial y} = \frac{2y - 2\beta x_0}{1 - \beta^2} = 0$$

a) Therefore $y = \beta x_0$.

For jointly Gaussian random variables

$$E[y|x] = E[y] + \frac{\text{Cov}(x, y)}{\text{Var}(x)} (x - E[x])$$

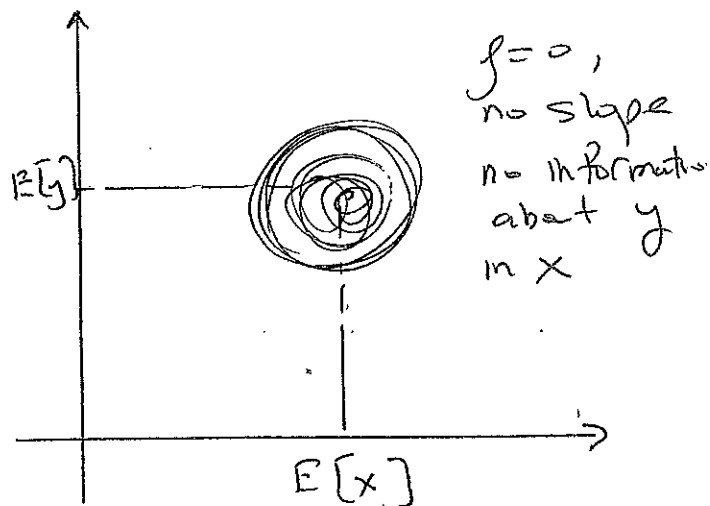
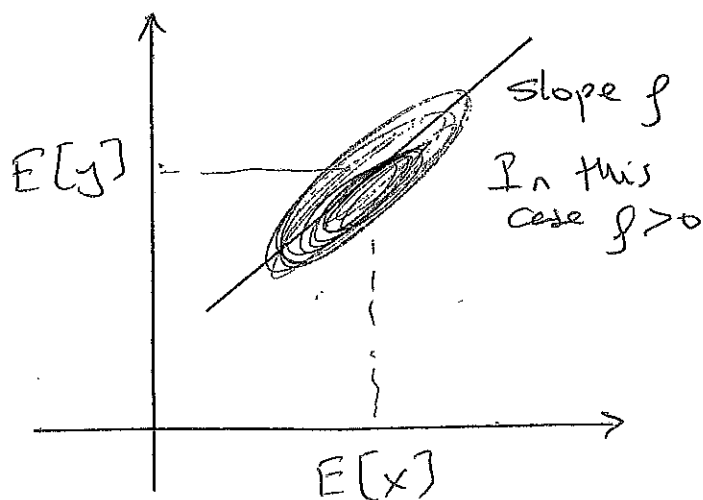
substituting for this problem

$$E[y|x] = 0 + \frac{\beta}{\text{variance}(x) \text{ from } C} (x - 0) = \beta x$$

b) Therefore $E[y|x=x_0] = \beta x_0$.

c) The mode & mean are the same because $p(y|x)$ is a Gaussian.

d) $\beta = 0 \Rightarrow E[y|x] = E[y] = 0$



④ 11.2 Gaussian Mixture

$$p(\theta|x) = \epsilon N(x, 1) + (1-\epsilon) N(-x, 1)$$

$$\underline{\text{MSE}} = E[\theta|x] = \epsilon x + (1-\epsilon)(-x)$$

$$\text{So } \epsilon = 0.5 \quad E[\theta|x] = 0$$

$$\epsilon = 3/4 \quad E[\theta|x] = \frac{3x}{4} - \frac{x}{4} = x/2$$

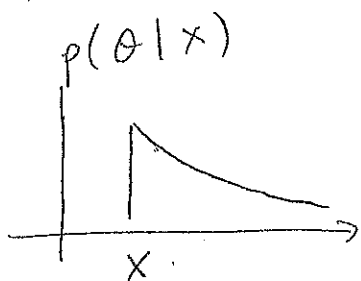
MAP For $\epsilon = 0.5$ MAP is not unique = either x or $-x$

For $\epsilon = 3/4$ MAP is x

11.3 exponential PDF

$$p(\theta|x) = \begin{cases} e^{-(\theta-x)} & , \theta > x \\ 0 & , \theta < x \end{cases}$$

$$\text{MMSE } E[\theta|x] = \int_x^\infty \theta e^{-(\theta-x)} d\theta$$



$$= e^x \left[-\theta e^{-\theta} - e^{-\theta} \right] \Big|_x^\infty$$
$$= e^x \left(x e^{-x} + e^{-x} \right) = x + 1$$



MAP estimator is simply x

$$\textcircled{5} \quad p_x(x[0], x[1] | A) = p_w(x[0]-A, x[1]-A | A)$$

$$= p_w(x[0]-A, x[1]-A) \quad \text{since } w \text{ and } A \text{ are independent}$$

$$= p_w(x[0]-A) p_w(x[1]-A) \quad \text{since } w[0] \text{ \& } w[1] \text{ are independent}$$

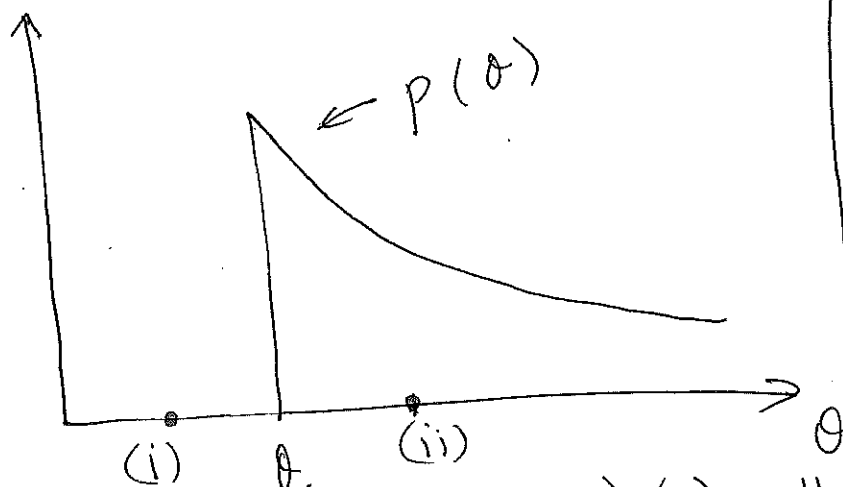
$$= p_x(x[0] | A) p_x(x[1] | A) \quad \checkmark$$

$$\textcircled{6} \quad \text{a) } p(\underline{x} | \theta) = \begin{cases} (1/\theta)^N & 0 \leq x[n] \leq \theta \text{ for all } n \\ 0, & \text{otherwise} \end{cases}$$

using conditional indep.

$$\text{Rewrite as } p(\underline{x} | \theta) = \begin{cases} (1/\theta)^N, & \min x[n] \geq 0 \\ & \max x[n] \leq \theta \\ 0, & \text{otherwise} \end{cases}$$

$$\text{b) } \hat{\theta}_{\text{MAP}} = \arg \max_{\theta} \underbrace{p(\underline{x} | \theta)}_{g(\theta)} p(\theta)$$



Combining the 2 cases:

$$\hat{\theta}_{\text{MAP}} = \max(\theta_0, \max x[n])$$

(i) $\max x[n] \leq \theta_0$ $p(\underline{x} | \theta) p(\theta)$ will be non-zero first at $\theta = \theta_0$ & since $p(\underline{x} | \theta)$ monotonically decreases with θ we must have $\hat{\theta}_{\text{MAP}} = \theta_0$

(ii) $\max x[n] > \theta_0$ $p(\underline{x} | \theta) p(\theta)$ will be first non-zero at $\theta = \max x[n]$ therefore $\hat{\theta}_{\text{MAP}} = \max x[n]$