

**THM:** Given  $n$  Bernoulli trials with probability of success for each trial being  $p$ , the probability,  $P(m \text{ of } n)$ , of exactly  $m$  successes in  $n$  trials approaches the probability density of  $x = m$  for a normal (i.e., gaussian) distribution with  $\mu = np$  and  $\sigma^2 = npq$ :

$$\text{As } n \rightarrow \infty, P(m \text{ of } n) \rightarrow f(x = m) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(m-\mu)^2 / 2\sigma^2}.$$

**PROOF:** We follow the general method of proof given in [1].

For Bernoulli trials we have the following value for  $P(m \text{ of } n)$ :

$$P(m \text{ of } n) = {}_n C_m \cdot p^m q^{n-m}$$

where  ${}_n C_m \equiv \frac{n!}{(n-m)!m!}$  is the combinatoric coefficient.

For the proof, we consider different values of  $n$ , and we will consider  $m$  to be a fixed number,  $k$ , of standard deviations from the mean as  $n$  increases.

$$m = \mu + k\sigma$$

**NOTE:** Although  $m$  is an integer, the method of proof allows  $k$  to have any real value.

We use Stirling's formula, [2], to approximate the factorials in  ${}_n C_m$ :

$$n! = \sqrt{2\pi} n^{n+\frac{1}{2}} e^{-n+\frac{\theta}{12n}}$$

where  $n > 0$  and  $0 < \theta < 1$ .

**NOTE:** Stirling's formula is related to the Stirling series expansion of the gamma function in powers of  $1/n$ , (see [3]). The Stirling series has the curious property that it produces very accurate approximations of the gamma functions with only a few terms—and actually diverges if all the terms are used.

Using Stirling's formula for the terms of  ${}_n C_m$ , yields the following expression:

$${}_n C_m = \frac{\sqrt{2\pi} n^{n+\frac{1}{2}} e^{-n+\frac{\theta_1}{12n}}}{\sqrt{2\pi} (n-m)^{n-m+\frac{1}{2}} e^{-n-m+\frac{\theta_2}{12(n-m)}} \sqrt{2\pi} m^{m+\frac{1}{2}} e^{-m+\frac{\theta_3}{12m}}}$$

As  $n$  becomes large, so do  $n - m$  and  $m$ , and the residual terms involving  $\theta_1$ ,  $\theta_2$ , and  $\theta_3$  become vanishingly small. Thus, we may eliminate the  $\theta$  terms and, after also canceling common factors of  $\sqrt{2\pi}$  and the exponentials of  $e$ , write the following expression:

$${}_n C_m \rightarrow \frac{n^{n+\frac{1}{2}}}{\sqrt{2\pi} (n-m)^{n-m+\frac{1}{2}} m^{m+\frac{1}{2}}} \text{ as } n \rightarrow \infty$$

If we split the  $n^n$  term into two pieces in the numerator, we can match up the exponents in the numerator and denominator:

$${}_n C_m \rightarrow \frac{n^{n-m} n^m \sqrt{n}}{\sqrt{2\pi} (n-m)^{n-m} \sqrt{n-m} \cdot m^m \sqrt{m}} \text{ as } n \rightarrow \infty$$

or

$${}_n C_m \rightarrow \frac{\sqrt{n}}{\sqrt{2\pi(n-m)m}} \left(\frac{n}{n-m}\right)^{n-m} \left(\frac{n}{m}\right)^m \text{ as } n \rightarrow \infty$$

Now we invert the terms being exponentiated and use the following formulas:

$$\frac{m}{n} = \frac{\mu + k\sigma}{n} = \frac{np + k\sigma}{n} = p \left(1 + \frac{k\sigma}{np}\right)$$

and

$$\frac{n-m}{n} = 1 - \frac{\mu + k\sigma}{n} = 1 - \frac{np + k\sigma}{n} = q \left(1 - \frac{k\sigma}{nq}\right)$$

Substituting these expressions yields the following equation:

$${}_n C_m \rightarrow \frac{\sqrt{n}}{\sqrt{2\pi(n-m)m}} \left(q \left(1 - \frac{k\sigma}{nq}\right)\right)^{-n+m} \left(p \left(1 + \frac{k\sigma}{np}\right)\right)^{-m} \text{ as } n \rightarrow \infty$$

The terms having  $n$  in their denominators will become small as  $n$  becomes large. Thus, we use an approximation that exploits this behavior:

$$\ln(1+x) \approx x - \frac{x^2}{2} \text{ for } x \text{ small}$$

or

$$(1+x)^r = e^{r \ln(1+x)} \approx e^{r \left( x - \frac{x^2}{2} \right)} \text{ for } x \text{ small (from Taylor series for } \ln)$$

Applying this identity to our formula for the combinatoric coefficient, we have the following expression:

$${}_n C_m \rightarrow \frac{\sqrt{n}}{\sqrt{2\pi(n-m)m}} q^{-n+m} p^{-m} e^{(-n+m) \left( -k \frac{\sigma}{nq} - \frac{k^2 \sigma^2}{2n^2 q^2} \right)} e^{-m \left( k \frac{\sigma}{np} - \frac{k^2 \sigma^2}{2n^2 p^2} \right)}$$

Using  $m = np + k\sigma$  and  $m - n = -nq + k\sigma$  we have

$${}_n C_m \rightarrow \frac{\sqrt{n}}{\sqrt{2\pi(n-m)m}} q^{-n+m} p^{-m} e^{(k\sigma - nq) \left( -k \frac{\sigma}{nq} - \frac{k^2 \sigma^2}{2n^2 q^2} \right) - (k\sigma + np) \left( k \frac{\sigma}{np} - \frac{k^2 \sigma^2}{2n^2 p^2} \right)}$$

If we consider just the exponent, we have the following calculation:

$$\begin{aligned} & -\frac{k^2 \sigma^2}{nq} - \frac{k^2 \sigma^2}{np} + k\sigma \left( \frac{k^2 \sigma^2}{2n^2 p^2} - \frac{k^2 \sigma^2}{2n^2 q^2} \right) + nq \frac{k^2 \sigma^2}{2n^2 q^2} + np \frac{k^2 \sigma^2}{2n^2 p^2} \\ & = -\frac{k^2 \sigma^2 p}{npq} - \frac{k^2 \sigma^2 q}{npq} + k\sigma \left( \frac{k^2 \sigma^2 q^2}{2n^2 p^2 q^2} - \frac{k^2 \sigma^2 p^2}{2n^2 p^2 q^2} \right) + (np + nq) \frac{k^2 \sigma^2}{2n^2 pq} \end{aligned}$$

Using  $\sigma^2 = npq$  the simplification of the exponent continues:

$$\begin{aligned} & = -k^2 p - k^2 q + k\sigma \left( \frac{k^2 q^2}{2\sigma^2} - \frac{k^2 p^2}{2\sigma^2} \right) + \frac{k^2}{2} \\ & = -k^2 + k \left( \frac{k^2 q^2}{2\sigma} - \frac{k^2 p^2}{2\sigma} \right) + \frac{k^2}{2} \\ & = -\frac{k^2}{2} + k \left( \frac{k^2 q^2}{2\sigma} - \frac{k^2 p^2}{2\sigma} \right) \end{aligned}$$

We observe that the second term is proportional to  $1/\sqrt{n}$  and vanishes as  $n$  becomes large. Dropping this term yields the following expression:

$${}_n C_m \rightarrow \frac{\sqrt{n}}{\sqrt{2\pi(n-m)m}} q^{-n+m} p^{-m} e^{-\frac{k^2}{2}} \text{ as } n \rightarrow \infty$$

If we now multiply by the probability,  $p^m q^{n-m}$  of one particular pattern of  $m$  successes occurring, we obtain the following expression:

$$P(m \text{ of } n) \rightarrow \frac{\sqrt{n}}{\sqrt{2\pi(n-m)m}} e^{-\frac{k^2}{2}} \text{ as } n \rightarrow \infty$$

We have the following simplification for the factor in front:

$$\sqrt{\frac{n}{2\pi(n-m)m}} = \sqrt{\frac{1}{2\pi(1-\frac{m}{n})m}} = \sqrt{\frac{1}{2\pi(1-\frac{np+k\sigma}{n})(np+k\sigma)}}$$

For  $n$  large,  $k\sigma$  is much smaller than  $n$ , leading to the following result:

$$\sqrt{\frac{n}{2\pi(n-m)m}} \approx \sqrt{\frac{1}{2\pi(1-\frac{np}{n})(np)}} = \sqrt{\frac{1}{2\pi qnp}} = \sqrt{\frac{1}{2\pi\sigma^2}}$$

With this substitution, and using  $k^2 = \frac{(m-\mu)^2}{\sigma^2}$  we complete our proof:

$$P(m \text{ of } n) \rightarrow \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{k^2}{2}} = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(m-\mu)^2/2\sigma^2} \text{ as } n \rightarrow \infty$$

- REF:** [1] Eugene Lukacs, *Probability and Mathematical Statistics, an Introduction*, New York, NY: Academic Press, 1972.
- [2] Milton Abramowitz and Irene A. Stegun, Eds., *Handbook of Mathematical Functions: National Bureau of Standards Applied Mathematics Series 55*, Washington, D.C.: Government Printing Office, 1972.
- [3] Carl M. Bender and Steven A. Orszag, *Advanced Mathematical Methods for Scientists and Engineers*, New York, NY: McGraw-Hill, 1978.